

Homework 6: finish by 6/12.

Reading: Notes: Chapter 4, 5.

Videos: 4.7, 5.1 - 5.2

Problem 6.1 ([Video 4.1 - 4.7](#), [Lecture Problem](#))

Let X be Uniform[1, 2]. Let Y given $X = x$ be Exponential(x). That is, let

$$f_{Y|X}(y|x) = \begin{cases} xe^{-xy} & y \geq 0 \\ 0 & y < 0 \end{cases}$$

- (a) Find the expected value of Y . (It is OK to leave your answer as an integral.)

Solution:

$$f_Y(y) = \int_{-\infty}^{\infty} f_{X,Y}(x, y) dx = \begin{cases} \int_1^2 xe^{-xy} dx & y \geq 0 \\ 0 & y < 0 \end{cases}$$

$$\begin{aligned} \mathbb{E}[Y] &= \int_{-\infty}^{\infty} y f_Y(y) dy = \int_0^{\infty} y \left(\int_1^2 xe^{-xy} dx \right) dy = \int_0^{\infty} \int_1^2 xye^{-xy} dx dy \\ &= \int_1^2 \int_0^{\infty} xye^{-xy} dy dx = \int_1^2 \left(\int_0^{\infty} ye^{-xy} dy \right) dx = \int_1^2 \frac{1}{x} dx = \ln(2) \end{aligned}$$

- (b) Find the conditional expected value $\mathbb{E}[Y|X = x]$ of Y given $X = x$. (This should be in closed form.)

Solution:

Recall that an Exponential(λ) random variable has expected value $\frac{1}{\lambda}$.
 Therefore, $\mathbb{E}[Y|X = x] = \frac{1}{x}$.

- (c) Using $\mathbb{E}[Y] = \mathbb{E}[\mathbb{E}[Y|X]]$ find the expected value of Y . (This should be in closed form.)

Solution:

$$\mathbb{E}[Y] = \mathbb{E}[\mathbb{E}[Y|X]] = \int_{-\infty}^{\infty} \mathbb{E}[Y|X = x] f_X(x) dx = \int_1^2 \frac{1}{x} dx = \ln(x) \Big|_1^2 = \ln(2)$$

- (d) Solve for $\mathbb{E}[XY]$. (It is OK to leave your answer as an integral.)

Solution:

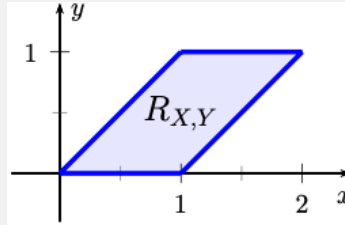
$$\mathbb{E}[XY] = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} xy f_{X,Y}(x, y) dx dy = \int_0^{\infty} \int_1^2 x^2 y e^{-xy} dx dy$$

Problem 6.2 (Video 4.6, 4.7, 5.1, 5.2, Lecture Problem)

Consider the joint PDF $f_{X,Y}(x, y) = \begin{cases} x & y \leq x \leq y+1 \text{ and } 0 \leq y \leq 1 \\ 0 & \text{otherwise} \end{cases}$

(a) Draw a sketch of the range $R_{X,Y}$.

Solution:



(b) What is the probability that $2Y$ is greater than X ?

Solution:

$$\begin{aligned} \mathbb{P}[2Y > X] &= \int_{-\infty}^{\infty} \int_{-\infty}^{2y} f_{X,Y} dx dy = \int_0^1 \int_y^{2y} x dx dy \\ &= \int_0^1 \left(\frac{1}{2} x^2 \right) \Big|_y^{2y} dy = \int_0^1 \frac{3}{2} y^2 dy = \frac{1}{2} y^3 \Big|_0^1 = \frac{1}{2} \end{aligned}$$

(c) Are X and Y independent? Explain.

Solution:

No, the range is not a rectangle, so the range cannot factor and thus the joint PDF does not factor either.

(d) Compute $E[X^2]$.

Solution:

$$\begin{aligned} \mathbb{E}[X^2] &= \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} x^2 f_{X,Y}(x, y) dx dy = \int_0^1 \int_y^{y+1} x^3 dx dy \\ &= \int_0^1 \left(\frac{1}{4} x^4 \right) \Big|_y^{y+1} dy = \int_0^1 \left(\frac{(y+1)^4}{4} - \frac{y^4}{4} \right) dy = \int_0^1 \left(\frac{4y^3 + 6y^2 + 4y + 1}{4} \right) dy \end{aligned}$$

$$= \left(\frac{y^4 + 2y^3 + 2y + y}{4} \right) \Big|_0^1 = \frac{3}{2}$$

(e) Compute $\text{Cov}[X, Y]$.

Solution:

We will use the formula $\text{Cov}[X, Y] = \mathbb{E}[XY] - \mathbb{E}[X]\mathbb{E}[Y]$. (For full credit, it suffices to give this formula accompanied by the integrals in blue below.)

$$\begin{aligned} \mathbb{E}[XY] &= \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} xy f_{X,Y} dx dy = \int_0^1 \int_y^{y+1} x^2 y dx dy \\ &= \int_0^1 \left(\frac{1}{3} x^3 \right) \Big|_y^{y+1} y dy = \int_0^1 \left(\frac{(y+1)^3}{3} - \frac{y^3}{3} \right) y dy = \int_0^1 \left(y^3 + y^2 + \frac{y}{3} \right) dy \\ &= \left(\frac{y^4}{4} + \frac{y^3}{3} + \frac{y^2}{6} \right) \Big|_0^1 = \frac{3}{4} \end{aligned}$$

$$\begin{aligned} \mathbb{E}[X] &= \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} x f_{X,Y} dx dy = \int_0^1 \int_y^{y+1} x^2 dx dy \\ &= \int_0^1 \left(\frac{1}{3} x^3 \right) \Big|_y^{y+1} dy = \int_0^1 \left(\frac{(y+1)^3}{3} - \frac{y^3}{3} \right) dy = \int_0^1 \left(\frac{3y^2 + 3y + 1}{3} \right) dy \\ &= \left(\frac{y^3}{3} + \frac{y^2}{2} + \frac{y}{3} \right) \Big|_0^1 = \frac{7}{6} \end{aligned}$$

$$\begin{aligned} \mathbb{E}[Y] &= \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} y f_{X,Y} dx dy = \int_0^1 \int_y^{y+1} xy dx dy \\ &= \int_0^1 \left(\frac{1}{2} x^2 \right) \Big|_y^{y+1} y dy = \int_0^1 \left(\frac{(y+1)^2}{2} - \frac{y^2}{2} \right) y dy = \int_0^1 \left(y^2 + \frac{y}{2} \right) dy \\ &= \left(\frac{y^3}{3} + \frac{y^2}{4} \right) \Big|_0^1 = \frac{7}{12} \end{aligned}$$

Putting this together, we have $\text{Cov}[X, Y] = \frac{3}{4} - \frac{7}{6} \cdot \frac{7}{12} = \frac{5}{72}$.

(f) Compute $\rho_{X,Y}$.

Solution:

We can use the formula $\rho_{X,Y} = \frac{\text{Cov}[X,Y]}{\sqrt{\text{Var}[X] \text{Var}[Y]}}$. To obtain the variances, we will use $\text{Var}[X] = \mathbb{E}[X^2] - (\mathbb{E}[X])^2$ and $\text{Var}[Y] = \mathbb{E}[Y^2] - (\mathbb{E}[Y])^2$. The only integral we are missing is

$$\begin{aligned} \mathbb{E}[Y^2] &= \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} y^2 f_{X,Y}(x, y) dx dy = \int_0^1 \int_y^{y+1} xy^2 dx dy \\ &= \int_0^1 \left(\frac{1}{2} x^2 \right) \Big|_y^{y+1} y^2 dy = \int_0^1 \left(\frac{(y+1)^2}{2} - \frac{y^2}{2} \right) y^2 dy = \int_0^1 \left(y^3 + \frac{y^2}{2} \right) dy \end{aligned}$$

$$= \left(\frac{y^4}{4} + \frac{y^3}{6} \right) \Big|_0^1 = \frac{5}{12}$$

Putting everything together, we have

$$\text{Var}[X] = \frac{3}{2} - \left(\frac{7}{6} \right)^2 = \frac{5}{36}$$

$$\text{Var}[Y] = \frac{5}{12} - \left(\frac{7}{12} \right)^2 = \frac{11}{144}$$

$$\rho_{X,Y} = \frac{\text{Cov}[X,Y]}{\sqrt{\text{Var}[X] \text{Var}[Y]}} = \frac{\frac{5}{72}}{\sqrt{\frac{5}{36} \cdot \frac{11}{144}}} = \sqrt{\frac{5}{11}} \approx 0.674$$

(g) Compute $f_{X|Y}(x|y)$.

Solution:

First, we will need the marginal PDF of Y .

$$f_Y(y) = \int_{-\infty}^{\infty} f_{X,Y}(x,y) dx = \begin{cases} \int_y^{y+1} x dx & 0 \leq y \leq 1 \\ 0 & \text{otherwise.} \end{cases} = \begin{cases} y + \frac{1}{2} & 0 \leq y \leq 1 \\ 0 & \text{otherwise.} \end{cases}$$

Now, we can use the conditional PDF formula:

$$f_{X|Y}(x|y) = \begin{cases} \frac{f_{X,Y}(x,y)}{f_Y(y)} & f_Y(y) > 0 \\ 0 & \text{otherwise} \end{cases} = \begin{cases} \frac{x}{y + \frac{1}{2}} & y \leq x \leq y + 1 \text{ and } 0 \leq y \leq 1 \\ 0 & \text{otherwise} \end{cases}$$

(h) Compute $\mathbb{E}[2X + 1|Y]$.

Solution:

By the linearity of expectation, $\mathbb{E}[2X + 1|Y] = 2\mathbb{E}[X|Y] + 1$. The conditional expected value of X given $Y = y$ is

$$\mathbb{E}[X|Y = y] = \int_{-\infty}^{\infty} x f_{X|Y}(x|y) dx = \int_y^{y+1} \frac{x^2}{y + \frac{1}{2}} dx = \frac{(y+1)^3 - y^3}{3(y + \frac{1}{2})} = \frac{2y^2 + 2y + \frac{2}{3}}{2y + 1}.$$

Therefore, we have that

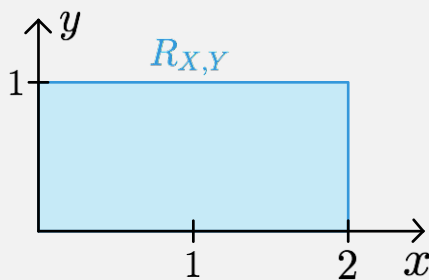
$$\mathbb{E}[2X + 1|Y] = 2 \frac{2Y^2 + 2Y + \frac{2}{3}}{2Y + 1} + 1 = \frac{4Y^2 + 4Y + \frac{4}{3} + 2Y + 1}{2Y + 1} = \frac{4Y^2 + 6Y + \frac{7}{3}}{2Y + 1}.$$

Problem 6.3 ([Video 4.6](#), [4.7](#), [5.1](#), [5.2](#))

Consider the following joint PDF $f_{X,Y}(x,y) = \begin{cases} \frac{x}{4} + \frac{y}{2} & 0 \leq x \leq 2, 0 \leq y \leq 1 \\ 0 & \text{otherwise.} \end{cases}$

(a) Draw a sketch of the range $R_{X,Y}$.

Solution:



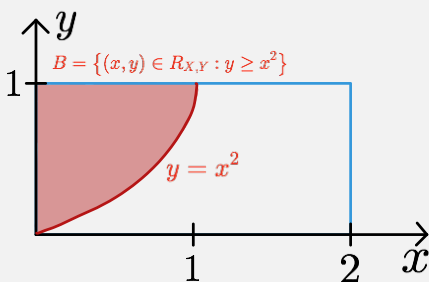
(b) Are X and Y independent?

Solution:

No, they are not independent. Even though the range is a rectangle, the joint PDF cannot be written as a product of the marginal PDFs (as we will see below).

(c) What is the probability that Y is greater than X^2 ?

Solution:



$$\begin{aligned}\mathbb{P}[Y > X^2] &= \int_{-\infty}^{\infty} \int_{x^2}^{\infty} f_{X,Y}(x, y) dy dx = \int_0^1 \int_{x^2}^1 \left(\frac{x}{4} + \frac{y}{2} \right) dy dx \\ &= \int_0^1 \left(\frac{xy}{4} + \frac{y^2}{4} \right) \Big|_{x^2}^1 dx = \int_0^1 \left(\frac{x}{4} + \frac{1}{4} - \frac{x^3}{4} - \frac{x^4}{4} \right) dx \\ &= \left(\frac{x^2}{8} + \frac{x}{4} - \frac{x^4}{16} - \frac{x^5}{20} \right) \Big|_0^1 = \frac{1}{8} + \frac{1}{4} - \frac{1}{16} - \frac{1}{20} = \frac{21}{80}\end{aligned}$$

(d) Determine the marginal PDFs $f_X(x)$ and $f_Y(y)$.

Solution:

$$\begin{aligned}
f_X(x) &= \int_{-\infty}^{\infty} f_{X,Y}(x,y) dy = \begin{cases} \int_0^1 \left(\frac{x}{4} + \frac{y}{2}\right) dy & 0 \leq x \leq 2 \\ 0 & \text{otherwise.} \end{cases} \\
&= \begin{cases} \left(\frac{xy}{4} + \frac{y^2}{4}\right)\Big|_0^1 & 0 \leq x \leq 2 \\ 0 & \text{otherwise.} \end{cases} = \begin{cases} \frac{x}{4} + \frac{1}{4} & 0 \leq x \leq 2 \\ 0 & \text{otherwise.} \end{cases} \\
f_Y(y) &= \int_{-\infty}^{\infty} f_{X,Y}(x,y) dx = \begin{cases} \int_0^2 \left(\frac{x}{4} + \frac{y}{2}\right) dx & 0 \leq y \leq 1 \\ 0 & \text{otherwise.} \end{cases} \\
&= \begin{cases} \left(\frac{x^2}{8} + \frac{xy}{2}\right)\Big|_0^2 & 0 \leq y \leq 1 \\ 0 & \text{otherwise.} \end{cases} = \begin{cases} y + \frac{1}{2} & 0 \leq y \leq 1 \\ 0 & \text{otherwise.} \end{cases}
\end{aligned}$$

(e) Calculate $\text{Cov}[X, Y]$.

Solution:

We will use the formula $\text{Cov}[X, Y] = \mathbb{E}[XY] - \mathbb{E}[X]\mathbb{E}[Y]$.

$$\begin{aligned}
\mathbb{E}[X] &= \int_{-\infty}^{\infty} x f_X(x) dx = \int_0^2 \left(\frac{x^2}{4} + \frac{x}{4}\right) dx = \left(\frac{x^3}{12} + \frac{x^2}{8}\right)\Big|_0^2 = \frac{8}{12} + \frac{4}{8} = \frac{7}{6} \\
\mathbb{E}[Y] &= \int_{-\infty}^{\infty} y f_Y(y) dy = \int_0^1 \left(y^2 + \frac{y}{2}\right) dy = \left(\frac{y^3}{3} + \frac{y^2}{4}\right)\Big|_0^1 = \frac{1}{3} + \frac{1}{4} = \frac{7}{12} \\
\mathbb{E}[XY] &= \int_0^2 \int_0^1 xy f_{X,Y}(x,y) dy dx = \int_0^2 \int_0^1 \left(\frac{x^2 y}{4} + \frac{xy^2}{2}\right) dy dx \\
&= \int_0^2 \left(\frac{x^2 y^2}{8} + \frac{xy^3}{6}\right)\Big|_0^1 dx = \int_0^2 \left(\frac{x^2}{8} + \frac{x}{6}\right) dx = \left(\frac{x^3}{24} + \frac{x^2}{12}\right)\Big|_0^2 = \frac{8}{24} + \frac{4}{12} = \frac{2}{3} \\
\text{Cov}[X, Y] &= \frac{2}{3} - \frac{7}{6} \cdot \frac{7}{12} = -\frac{1}{72}
\end{aligned}$$

(f) Calculate $\rho_{X,Y}$.

Solution:

To compute the correlation coefficient, $\rho_{X,Y} = \frac{\text{Cov}[X,Y]}{\sqrt{\text{Var}[X]\text{Var}[Y]}}$, we first need to compute the variances. Using the formulas $\text{Var}[X] = \mathbb{E}[X^2] - (\mathbb{E}[X])^2$ and $\text{Var}[Y] = \mathbb{E}[Y^2] - (\mathbb{E}[Y])^2$, we see that we just need to calculate the second moments.

$$\begin{aligned}
\mathbb{E}[X^2] &= \int_{-\infty}^{\infty} x^2 f_X(x) dx = \int_0^2 \left(\frac{x^3}{4} + \frac{x^2}{4}\right) dx = \left(\frac{x^4}{16} + \frac{x^3}{12}\right)\Big|_0^2 = \frac{16}{16} + \frac{8}{12} = \frac{5}{3} \\
\mathbb{E}[Y^2] &= \int_{-\infty}^{\infty} y^2 f_Y(y) dy = \int_0^1 \left(y^3 + \frac{y^2}{2}\right) dy = \left(\frac{y^4}{4} + \frac{y^3}{6}\right)\Big|_0^1 = \frac{1}{4} + \frac{1}{6} = \frac{5}{12}
\end{aligned}$$

Putting things together, we get

$$\begin{aligned}\text{Var}[X] &= \frac{5}{3} - \left(\frac{7}{6}\right)^2 = \frac{11}{36} \\ \text{Var}[Y] &= \frac{5}{12} - \left(\frac{7}{12}\right)^2 = \frac{11}{144} \\ \rho_{X,Y} &= \frac{-\frac{1}{72}}{\sqrt{\frac{11}{36} \cdot \frac{11}{144}}} = -\frac{6 \cdot 12}{11 \cdot 72} = -\frac{1}{11}\end{aligned}$$

- (g) Determine the conditional PDF $f_{X|Y}(x|y)$.

Solution:

$$f_{X|Y}(x|y) = \begin{cases} \frac{f_{X,Y}(x,y)}{f_Y(y)} & f_Y(y) > 0 \\ 0 & \text{otherwise.} \end{cases} = \begin{cases} \frac{\frac{x}{4} + \frac{y}{2}}{y + \frac{1}{2}} & 0 \leq x \leq 2, 0 \leq y \leq 1 \\ 0 & \text{otherwise.} \end{cases}$$

- (h) Calculate the conditional expected value $\mathbb{E}[X|Y = y]$.

Solution:

$$\begin{aligned}\mathbb{E}[X|Y = y] &= \int_{-\infty}^{\infty} x f_{X|Y}(x|y) dx \\ &= \int_0^2 x \frac{\frac{x}{4} + \frac{y}{2}}{y + \frac{1}{2}} dx = \frac{1}{y + \frac{1}{2}} \int_0^2 \left(\frac{x^2}{4} + \frac{xy}{2} \right) dx = \frac{1}{y + \frac{1}{2}} \left(\frac{x^3}{12} + \frac{x^2 y}{4} \right) \Big|_0^2 = \frac{y + \frac{2}{3}}{y + \frac{1}{2}}\end{aligned}$$

- (i) Using the iterated expectation formula $\mathbb{E}[\mathbb{E}[X|Y]] = \mathbb{E}[X]$ and your answer from part (h), determine $\mathbb{E}[X]$. Compare with the answer you found from the marginals as part of your calculations for part (e).

Solution:

From the iterated expectation formula,

$$\begin{aligned}\mathbb{E}[X] &= \mathbb{E}[\mathbb{E}[X|Y]] = \int_{-\infty}^{\infty} \mathbb{E}[X|Y = y] f_Y(y) dy \\ &= \int_0^1 \frac{y + \frac{2}{3}}{y + \frac{1}{2}} \left(y + \frac{1}{2} \right) dy = \int_0^1 \left(y + \frac{2}{3} \right) dy = \left(\frac{1}{2} y^2 + \frac{2y}{3} \right) \Big|_0^1 = \frac{1}{2} + \frac{2}{3} = \frac{7}{6},\end{aligned}$$

which agrees with our calculations from part (e).

Problem 6.4 (Video 4.5, 4.7, Quick Calculations)

Parts (a) - (d): For each of the scenarios below, determine if X and Y are independent, and give a justification. (You should be able to do this without any long calculations or integration.)

(a) $f_{X,Y}(x,y) = \begin{cases} 2 & 0 \leq x \leq y \leq 1 \\ 0 & \text{otherwise} \end{cases}$

Solution:

Not independent since the range is not a rectangle.

- (b) Y given X is Geometric($\frac{x+1}{3}$) and X is Bernoulli($1/2$).

Solution:

Not independent since $P_{Y|X}(y|x)$ depends on x and thus cannot be equal to the marginal $P_Y(y)$.

$$(c) f_{X,Y}(x, y) = \begin{cases} \frac{9}{38}x^2y^2 & -1 \leq x \leq 1, 2 \leq y \leq 3 \\ 0 & \text{otherwise} \end{cases}$$

Solution:

Yes, since the range is a rectangle *and* the function will clearly factor into components of the form c_1x^2 and c_2y^2 for some constants c_1 and c_2 . Therefore, we can conclude that $f_{X,Y}(x, y) = f_X(x)f_Y(y)$.

(d)

$P_{XY}(x, y)$		y	
		0	1
x	0	1/15	4/15
	1	2/15	8/15

Solution:

Yes, we can solve for the marginals $P_X(x) = \begin{cases} \frac{1}{3} & x = 0 \\ \frac{2}{3} & x = 1 \\ 0 & \text{otherwise} \end{cases}$ and $P_Y(y) = \begin{cases} \frac{1}{5} & y = 0 \\ \frac{4}{5} & y = 1 \\ 0 & \text{otherwise} \end{cases}$ and then verify that $P_{X,Y}(x, y) = P_X(x)P_Y(y)$.

$$(e) f_{X,Y}(x, y) = \begin{cases} \frac{1}{2} & 0 \leq x \leq 1, 0 \leq y \leq 1 \\ \frac{1}{2} & -1 \leq x \leq 0, -1 \leq y \leq 0 \\ 0 & \text{otherwise} \end{cases}$$

Solution:

No, the range is not a rectangle, and thus the joint PDF will not factor into the product of the marginal PDFs.

Parts (e) - (f): For each of the scenarios below, determine the requested quantities. (You should be able to do this without any long calculations or integration.)

- (e) Let X be Poisson(3) and Y given $X = x$ be Binomial($x, \frac{1}{3}$). Calculate $\mathbb{E}[Y|X = x]$ and $\mathbb{E}[Y]$.

Solution:

For a Binomial(n, p) random variable, the expected value is np . Therefore, $\mathbb{E}[Y|X = x] = \frac{x}{3}$. Using the iterated expectation formula,

$$\mathbb{E}[Y] = \mathbb{E}[\mathbb{E}[Y|X]] = \mathbb{E}\left[\frac{X}{3}\right] = \frac{1}{3}\mathbb{E}[X] = \frac{1}{3} \cdot 3 = 1 .$$

- (f) Let X be Uniform($\frac{1}{2}, \frac{5}{2}$). Let Y given $X = x$ be Exponential($\frac{1}{2x}$). Calculate $\mathbb{E}[Y|X = x]$ and $\mathbb{E}[Y]$.

Solution:

First, recall that for an Exponential(λ) random variable, the mean is $\frac{1}{\lambda}$ and the variance is $\frac{1}{\lambda^2}$. Therefore,

$$\mathbb{E}[Y|X = x] = \frac{1}{\frac{1}{2x}} = 2x .$$

Using the iterated expectation formula,

$$\mathbb{E}[Y] = \mathbb{E}[\mathbb{E}[Y|X]] = \mathbb{E}[2X] = 2 \cdot \frac{\frac{5}{2} + \frac{1}{2}}{2} = 2 \cdot 3 = 6 .$$