Exam 2 Solutions

Problem 1

16 points

Let X be a continuous random variable with the PDF $f_X(x) = \frac{1}{2}e^{-|x+1|}$ called the double exponential or Laplace distribution. The range of X is \mathbb{R} , i.e., the set of all real numbers.

(a) (4 pts) Compute $\mathbb{E}[X]$. Your answer can be an integral, but you can also exploit symmetry to get an exact expression.

Solution: $\mathbb{E}[X] = -1.$

Method 1 using symmetry: The PDF $f_X(x)$ is symmetric around -1: $f_X(-1+t) = f_X(-1-t)$ for all t. Thus, $\mathbb{E}[X] = -1$. Method 2 using integrals: $\mathbb{E}[X] = \int_{-\infty}^{\infty} x f_X(x) dx = \int_{-\infty}^{\infty} \frac{xe^{-|x+1|}}{2} dx = -1$

(b) (4 pts) Let $A = \{X > 0\}$. Compute expressions of $\mathbb{P}[X \in A]$ and the conditional PDF $f_{X|A}(x)$ (as a case-by-case formula). The answers are simple expressions, but they can be left in terms of integrals.

Solution:
$$\mathbb{P}[A] = \frac{1}{2e}, \ f_{X|A}(x) = \text{PDF of an Exponential(1) RV.}$$

 $\mathbb{P}[A] = \int_0^\infty f_X(x) dx = \int_0^\infty \frac{1}{2} e^{-|x+1|} dx = \int_0^\infty \frac{1}{2} e^{-(x+1)} dx = \frac{e^{-1}}{2} \int_0^\infty e^{-x} dx = \frac{e^{-1}}{2} = \frac{1}{2e}$

$$f_{X|A}(x) = \begin{cases} \frac{f_X(x)}{\mathbb{P}[X \in A]} & x \in A\\ 0 & \text{otherwise} \end{cases} = \begin{cases} ee^{-|x+1|} & x > 0\\ 0 & \text{otherwise} \end{cases} = \begin{cases} ee^{-(x+1)} & x > 0\\ 0 & \text{otherwise} \end{cases} = \begin{cases} e^{-x} & x > 0\\ 0 & \text{otherwise} \end{cases}$$

Observe that the conditional PDF of X when given that X > 0, is that of an Exponential(1) RV.

(c) (4 pts) Compute $\mathbb{P}[X < 2|X > 0]$. The answer is a simple expression, but it can be left in terms of integrals.

Solution: $1 - e^{-2}$

Method 1 using conditional probability definition:

$$\mathbb{P}[X < 2|X > 0] = \frac{\mathbb{P}[0 < X < 2]}{\mathbb{P}[X > 0]} = \frac{\int_0^2 0.5e^{-|x+1|} dx}{\int_0^\infty 0.5e^{-|x+1|} dx} = \frac{\int_0^2 0.5e^{-(x+1)} dx}{\int_0^\infty 0.5e^{-(x+1)} dx} = \frac{\int_0^2 e^{-x} dx}{\int_0^\infty e^{-x} dx} = \frac{1 - e^{-2}}{1 - 0} = 1 - e^{-2}$$

Method 2 using conditional PDF from part (b)

$$\mathbb{P}[X < 2|X > 0] = \mathbb{P}[X < 2|X \in A] = \int_{-\infty}^{2} f_{X|A}(x)dx = \int_{0}^{2} e^{-x}dx = 1 - e^{-2x}dx$$

Method 3 using properties of Exponential RVs: Since the PDF of X given X > 0 is that of an Exponential(1) RV, $\mathbb{P}[X < 2|X > 0] = 1 - e^{-2}$.

(d) (4 pts) Compute $\mathbb{E}[X|X > 0]$. The answer is a simple expression, but it can be left in terms of integrals.

Solution: 1

Method 1 using conditional expectation definition:

$$\mathbb{E}[X|X>0] = \mathbb{E}[X|X\in A] = \int_{-\infty}^{\infty} x f_{X|A}(x) dx = \int_{0}^{\infty} x e^{-x} dx = 1$$

Method 2 using properties of Exponential RVs: Since the PDF of X given X > 0 is that of an Exponential(1) RV, $\mathbb{E}[X|X > 0] = 1$.

Problem 2

16 points

Consider the pair of discrete random variables X, Y with joint PMF described below:

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$P_{X,Y}(x,y)$		-1	0	1	2
	1	1/12	1/12	1/12	1/12
x	2	0	1/6	1/6	0
	3	1/6	0	0	1/6

(a) (4 pts) Compute $\mathbb{P}[XY > 1]$.

Solution: $\mathbb{P}[XY > 1] = 5/12$

From the table, there are three non-zero entries with XY > 1: they are (1,2), (2,1) and (3,2). Their probabilities are 1/12, 1/6, and 1/6, so $\mathbb{P}[XY > 1] = 5/12$.

(b) (4 pts) Compute $\mathbb{E}[X]$ and $\mathbb{E}[Y]$.

Solution: $\mathbb{E}[X] = 2, \mathbb{E}[Y] = 0.5$

First, compute the marginal PMFs $P_X(x)$ and $P_Y(y)$ by summing along the rows and columns, respectively. This shows that

$$P_X(1) = P_X(2) = P_X(3) = 1/3 \Rightarrow X \quad \sim \text{Uniform}(1,3) \Rightarrow \mathbb{E}[X] = (1+3)/2 = 2.$$

$$P_Y(-1) = P_Y(0) = P_Y(1) = P_Y(2) = 1/4 \Rightarrow Y \quad \sim \text{Uniform}(-1,2) \Rightarrow \mathbb{E}[Y] = (-1+2)/2 = 0.5.$$

(c) (4 pts) Compute $\operatorname{Var}[Y|X=3]$.

Solution: Var[Y|X = 3] = 2.25

When X = 3, $Y \sim 3$ Bernoulli(1/2) - 1. Since the variance of Bernoulli(1/2) is 1/4, $Var[Y|X = 3] = (3)^2 * (1/4) = 9/4 = 2.25$.

(d) (4 pts) Compute $\rho_{X,Y}$.

Solution: $\rho_{X,Y} = 0$ Method 1: Using $Cov[X,Y] = \mathbb{E}[XY] - \mathbb{E}[X]\mathbb{E}[Y]$, we need $\mathbb{E}[XY]$. Using the full joint PMF,

$$\mathbb{E}[XY] = \frac{1(1)(-1) + 1(1)(1) + 1(1)(2) + 2(2)(1) + 2(3)(-1) + 2(3)(2)}{12} = 1$$

Hence Cov[X, Y] = 1 - (2)(0.5) = 0 and $\rho_{X,Y} = 0$ (you don't need to know the variances of X and Y).

Method 2: Observe that each value of y, $P_{X,Y}[0.5 - t, y] = P_{X,Y}[0.5 + t, y]$, i.e., the joint PMF is symmetric with respect to the vertical line x = 0.5. Therefore X and Y are uncorrelated and $Cov[X, y] = \rho_{X,Y} = 0$.

Problem 3

16 points

You win a hundred dollars in the lottery! Feeling generous, you first give an amount $X \sim \text{Uniform}(0, 100)$ of your winnings to one of your friends, and then give an amount $Y \sim \text{Uniform}(0, 100-X)$ to another friend. Both X and Y are continuous random variables.

(a) (4 pts) Are X and Y independent? Why or why not?

Solution: No. Because the range of Y given X = x is 100 - x which depends on the value of x.

(b) (4 pts) Compute the joint PDF $f_{X,Y}(x,y)$ and clearly state its range.

Solution:

$$f_X(x) = \begin{cases} \frac{1}{100} & 0 \le x \le 100\\ 0 & \text{otherwise} \end{cases}, \quad f_{Y|X}(y|x) = \begin{cases} \frac{1}{100-x} & 0 \le y \le 100 - x\\ 0 & \text{otherwise} \end{cases}$$
$$\Rightarrow f_{X,Y}(x,y) = f_X(x)f_{Y|X}(y|x) = \begin{cases} \frac{1}{100(100-x)} & 0 \le x \le 100, 0 \le y \le 100 - x\\ 0 & \text{otherwise} \end{cases}$$

Range $R_{X,Y} = \{(x, y) : 0 \le x \le 100, 0 \le y \le 100 - x\}.$

(c) (4 pts) Compute $\mathbb{E}[Y|X=20]$. For full credit, you must provide an exact numerical value.

Solution: $\mathbb{E}[Y|X=20] = 40$ $Y|X=20 \sim \text{Uniform}(0, 100-20) \Rightarrow \mathbb{E}[Y|X=20] = \frac{0+(100-20)}{2} = 40.$

(d) (4 pts) Compute $\mathbb{E}[Y]$. For full credit, you must provide an exact numerical value.

Solution:
$$\mathbb{E}[Y] = 25$$

 $\mathbb{E}[X] = \frac{0+100}{2} = 50.$
 $Y|X = x \sim \text{Uniform}(0, 100 - x) \Rightarrow \mathbb{E}[Y|X] = \frac{0+(100-X)}{2} = 50 - \frac{X}{2}.$
By the law of total probability, $\mathbb{E}[Y] = \mathbb{E}[\mathbb{E}[Y|X]] = \mathbb{E}[50 - 0.5X] = 50 - 0.5\mathbb{E}[X] = 50 - 25 = 25.$

16 points

Problem 4

Let X_1 and X_2 be independent standard Gaussian (zero mean, unit variance) RVs and $\underbrace{\begin{bmatrix} Y_1 \\ Y_2 \end{bmatrix}}_{\underline{Y}} = \underbrace{\begin{bmatrix} 1 & 1 \\ 1 & -1 \end{bmatrix}}_{A} \underbrace{\begin{bmatrix} X_1 \\ X_2 \end{bmatrix}}_{\underline{X}}$

(a) (4 pts) Compute exact numerical values of the 2×1 mean vector $\mu_{\underline{Y}} = \mathbb{E}[\underline{Y}]$ and the 2×2 covariance matrix $\Sigma_{\underline{Y}} = \mathsf{Cov}[\underline{Y}]$.

Solution: $\mu_{\underline{Y}} = \begin{bmatrix} 0 \\ 0 \end{bmatrix}, \Sigma_{\underline{Y}} = \begin{bmatrix} 2 & 0 \\ 0 & 2 \end{bmatrix}$

Since X_1, X_2 are independent standard Gaussian RVs,

$$\mu_{\underline{X}} = \begin{bmatrix} 0\\ 0 \end{bmatrix} = \underline{0}$$
, the 2 x 1 zero vector and $\Sigma_{\underline{X}} = \begin{bmatrix} 1 & 0\\ 0 & 1 \end{bmatrix} = I_2$, the 2 x 2 identity matrix

Therefore,

$$\mu_{\underline{Y}} = A \ \mu_{\underline{X}} = A \ \underline{0} = \begin{bmatrix} 0 \\ 0 \end{bmatrix}, \text{ and } \Sigma_{\underline{Y}} = A \Sigma_{\underline{X}} A^{\top} = A I_2 A^{\top} = A A^{\top} = \begin{bmatrix} 1 & 1 \\ 1 & -1 \end{bmatrix} \begin{bmatrix} 1 & 1 \\ 1 & -1 \end{bmatrix} = \begin{bmatrix} 2 & 0 \\ 0 & 2 \end{bmatrix}$$

(b) (4 pts) Compute $\mathbb{E}[Y_1|X_1]$ as a function of the RV X_1 .

Solution: $\mathbb{E}[Y_1|X_1] = X_1$

Method 1: Since X_1, X_2 are independent standard Gaussian RVs, $\mathbb{E}[Y_1|X_1] = \mathbb{E}[X_1 + X_2|X_1] = \mathbb{E}[X_1|X_1] + \mathbb{E}[X_2|X_1] = X_1 + \mathbb{E}[X_2] = X_1 + 0 = X_1$.

Method 2: First note that since X_1, X_2 are independent standard Gaussian RVs, they are uncorrelated and have unit variance, i.e., $\mathsf{Cov}[X_1, X_2] = 0$ and $\mathsf{Var}[X_1] = \mathsf{Cov}[X_1, X_1] = 1 = \mathsf{Var}[X_2]$.

Since $Y_1 = X_1 + X_2$ and $X_1 = X_1 + 0 \times X_2$ are linear functions of independent standard Gaussian RVs X_1, X_2 , it follows that X_1, Y_1 are jointly Gaussian and therefore $Y_1|X_1 = x_1$ is a Gaussian RV with mean given by

$$\mathbb{E}[Y_1|X_1 = x_1] = \mu_{Y_1} + \frac{\mathsf{Cov}[Y_1, X_1]}{\mathsf{Var}[Y_1]}(x_1 - \mu_{X_1}) = 0 + \frac{\mathsf{Cov}[X_1 + X_2, X_1]}{2}(x_1 - 0) = \frac{\mathsf{Cov}[X_1, X_1] + \mathsf{Cov}[X_2, X_1]}{2}x_1 = \frac{x_1}{2}.$$

Therefore, $\mathbb{E}[Y_1|X_1] = X_1/2.$

(c) (4 pts) Compute $\mathbb{P}[Y_1 \leq b | X_1 = a]$ in terms of a, b and the standard Gaussian CDF $\Phi(\cdot)$.

Solution: $\mathbb{P}[Y_1 \le b | X_1 = a] = \Phi(b - 0.5a)$

From solution method 2 of part (b), $Y_1|X_1 = a$ is a Gaussian RV with mean $E[Y_1|X_1 = a] = 0.5a$. We also have

$$\begin{aligned} \mathsf{Var}[Y_1|X_1 = a] = \mathsf{Var}[Y_1] - \frac{(\mathsf{Cov}[Y_1, X_1])^2}{\mathsf{Var}[X_1]} &= 2 - \frac{1^2}{1} = 1, \\ \end{aligned}$$
erefore $\mathbb{P}[Y_1 \le b | X_1 = a] = \Phi\left(\frac{b - \mathbb{E}[Y_1|X_1 = a]}{\sqrt{\mathsf{Var}[Y_1|X_1 = a]}}\right) = \Phi\left(\frac{b - 0.5a}{\sqrt{1}}\right) = \Phi(b - 0.5a). \end{aligned}$

(d) (4 pts) Compute the exact numerical value of $\mathbb{E}[Y_1^2 Y_2^2]$.

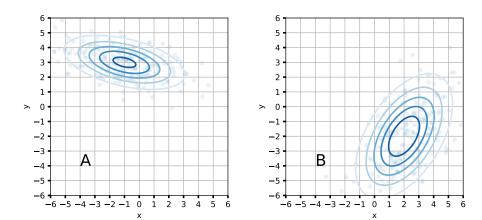
Solution: $\mathbb{E}[Y_1^2Y_2^2] = 4$ Since Y_1, Y_2 are linear functions of independent standard Gaussian RVs X_1, X_2 , they are jointly Gaussian. From part (a), $\mathsf{Cov}[Y_1, Y_2] = 0 \Rightarrow Y_1, Y_2$ are uncorrelated and since they are jointly Gaussian, they are independent RVs. Moreover, from part (a) $\mathbb{E}[Y_1] = \mathbb{E}[Y_2] = 0$ and therefore $\mathbb{E}[Y_1^2] = \mathsf{Var}[Y_1] = 2 = \mathsf{Var}[Y_2] = \mathbb{E}[Y_2^2]$. Therefore, $\mathbb{E}[Y_1^2Y_2^2] = \mathbb{E}[Y_1^2]\mathbb{E}[Y_2^2] = 2 \times 2 = 4$.

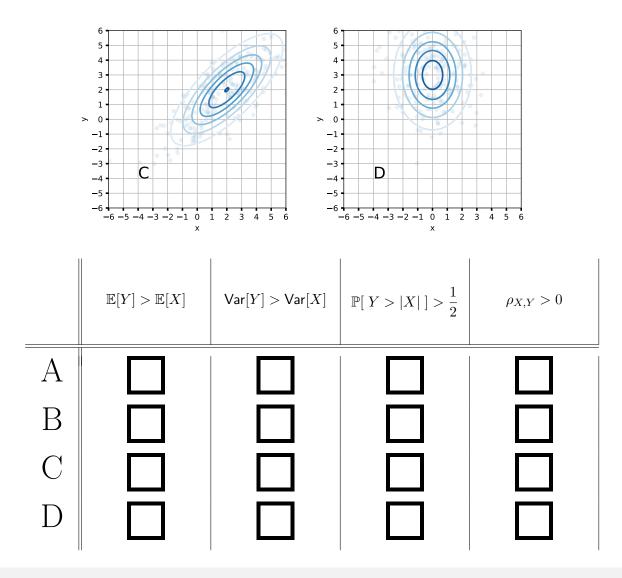
Problem 5

The

16 points

The table below depicts four jointly Gaussian PDFs via contour plots. In each case, the expectations, variances, and covariances are small integer values between -4 and 4. Put a checkmark in the boxes in each column that you think are true for that contour plot. No justifications are needed and there may be multiple boxes checked per row and/or column.



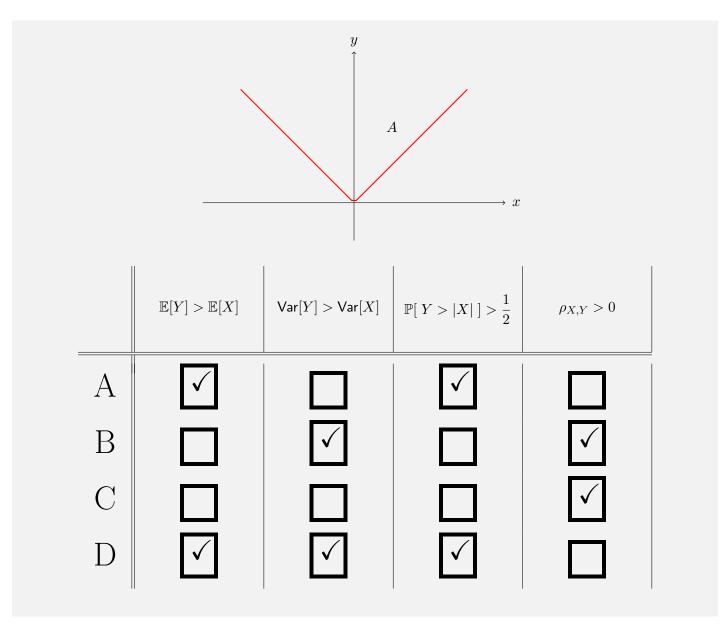


Solution:

The actual distributions are all jointly Gaussian, with the following parameters:

	$\mathbb{E}[X]$	$\mathbb{E}[Y]$	Var[X]	Var[Y]	Cov[X,Y]
A	-1	3	5	1	-1
В	2	-2	3	5	2
C	2	2	4	4	3
D	0	3	2	4	0

Note that the event $\{Y > |X|\} = \{X \ge 0, Y > X\} \cup \{X < 0, Y > -X\} = \{(X, Y) \in A\}$ where A is the V-shaped triangular region above the red curve y = |x| shown in the figure below.



Problem 6

8 points

Complete the following quick calculations. For full credit, you must work out a simplified, numerical answer for each requested quantity in this problem. The solutions do not require integration.

(a) (2pts) Let X be a standard Gaussian RV. Compute $\mathbb{P}[|X| > 1|X < 2]$ in terms of the standard Gaussian CDF $\Phi(\cdot)$.

Solution:
$$\begin{split} \mathbb{P}[|X| > 1 | X < 2] &= \frac{\phi(-1) + \phi(2) - \phi(1)}{\phi(2)} \\ \mathbb{P}[|X| > 1 | X < 2] &= \frac{\mathbb{P}[\{|X| > 1\} \cap \{X < 2\}]}{\mathbb{P}[X < 2]} = \frac{\mathbb{P}[\{X < -1\}] + \mathbb{P}[\{X < 2\}] - \mathbb{P}[\{X < 1\}]}{\mathbb{P}[X < 2]} \\ &= \frac{\phi(-1) + \phi(2) - \phi(1)}{\phi(2)} \end{split}$$

(b) (2pts) Let X be a continuous Uniform (-1, 1) RV. Compute $\mathbb{E}[X|X^2 > 0.25]$.

Solution: $\mathbb{E}[X|X^2 > 0.25] = 0$

The PDF of X conditioned on the event $A = \{X^2 > 0.25\} = \{-1 \le X < -0.5\} \cup \{0.5 < X \le 1\}$, is uniform over $[-1, -0.5) \cup (0.5, 1]$ which is symmetric about 0. Thus, $\mathbb{E}[X|X^2 > 0.25] = 0$.

(c) (2pts) Let X be continuous Uniform(-1,1) RV and let RV Y given X = x be Exponential $\left(\frac{1}{1+x^2}\right)$. Compute $\mathbb{E}[Y]$.

Solution: $\mathbb{E}[Y] = 4/3 \approx 1.33$

First note that since X is a continuous Uniform(-1,1) RV, $\mathbb{E}[X] = \frac{-1+1}{2} = 0$ and therefore $\mathbb{E}[X^2] = \operatorname{Var}[X] = \frac{(1-(-1))^2}{12} = \frac{4}{12} = \frac{1}{3}$. Next, since Y given X = x is an $\operatorname{Exponential}\left(\frac{1}{1+x^2}\right)$ RV, $\mathbb{E}[Y|X = x] = \frac{1}{\frac{1}{1+x^2}} = 1 + x^2$. Thus, $\mathbb{E}[Y|X] = 1 + X^2$. Finally, by the law of total expectation, $\mathbb{E}[Y] = \mathbb{E}[\mathbb{E}[Y|X]] = \mathbb{E}[1 + X^2] = 1 + \frac{1}{3} = 4/3 \approx 1.33$.

(d) (2pts) Compute $\mathbb{E}[(X+Y)^2]$ if X is Exponential(1), Y a standard Gaussian, and $\rho_{X,Y} = -0.5$.

Solution: $\mathbb{E}[(X+Y)^2] = 2$

Since X be Exponential(1), we have $\mu_X = \frac{1}{1} = 1$ and $\operatorname{Var}[X] = \frac{1}{1^2} = 1 \Rightarrow \mathbb{E}[X^2] = \operatorname{Var}[X] + \mu_X^2 = 2$. Since Y a standard Gaussian, we have $\mu_Y = 0$ and $\operatorname{Var}[Y] = 1 = \mathbb{E}[Y^2]$. We also have $\operatorname{Cov}[X,Y] = \rho_{X,Y} \sqrt{\operatorname{Var}[X]} \sqrt{\operatorname{Var}[Y]} = -0.5 \times 1 \times 1 = -0.5$. Therefore, $\mathbb{E}[XY] = \operatorname{Cov}[X,Y] + \mu_X \mu_Y = \operatorname{Cov}[X,Y] = -0.5$ Finally, $\mathbb{E}[(X+Y)^2] = \mathbb{E}[X^2] + \mathbb{E}[Y^2] + 2\mathbb{E}[XY] = 2 + 1 + 2(-0.5) = 2 + 1 - 1 = 2$.

Problem 7

8 points

For each of the following parts, indicate whether the statement is always true or it can be false by clearly writing T (for True) or F (for False) in **the box next to the question**. Full credit will be given for selecting the correct logical value (even with no explanation). Briefly explain your reasoning in the space provided for partial credit. Diagrams are welcome.

(a)
$$(2pts)$$
 If X is a continuous uniform RV with $\mathbb{E}[X] = 1$ and $\operatorname{Var}[X] = \frac{1}{3}$, then $\mathbb{P}[X < 0] = 0$.

Solution: True Let X be Uniform(a, b). Then $(a+b)/2 = 1 \Rightarrow (a+b) = 2$ and $(b-a)^2/12 = 1/3 \Rightarrow (b-a) = 2$. This implies that a = 0 and b = 2 and the range of X is [0, 2]. Thus, $\mathbb{P}[X < 0] = 0$.

(b) (2pts) If X is a continuous Uniform(0,3) RV, then $Y = X^2$ is a continuous Uniform(0,9) RV.

Solution: False $\mathbb{P}[Y \leq 1] = \mathbb{P}[X \leq 1] = 1/3$. If Y was a continuous Uniform(0,9) RV, then $\mathbb{P}[Y \leq 1] = 1/9$.

(c) (2pts) If
$$\operatorname{Var}[X] = 0.01$$
 and $\operatorname{Var}[Y] = 0.04$ then $\operatorname{Cov}[X, Y]$ can be 0.03.

Solution: False For the given values, $\rho_{X,Y} = \frac{\mathsf{Cov}[X,Y]}{\sqrt{\mathsf{Var}[X]\mathsf{Var}[Y]}} = \frac{0.03}{\sqrt{0.01 \times 0.04}} = \frac{0.03}{0.1 \times 0.2} = 1.5$ which is impossible, since $|\rho_{X,Y}| \le 1$.

(d) (2pts) If
$$\operatorname{Var}[-2X+3Y] = 4\operatorname{Var}[X] + 9\operatorname{Var}[Y]$$
, then $\operatorname{Var}[X-2Y] = \operatorname{Var}[X] + 4\operatorname{Var}[Y]$.

Solution: True $Var[aX + bY] = a^2 Var[X] + b^2 Var[Y] + 2abCov[X, Y]$. So if Var[-2X + 3Y] = 4Var[X] + 9Var[Y] then Cov[X, Y] = 0 and therefore Var[X - 2Y] = Var[X] + 4Var[Y].

4 points

Problem 8

X and Y are RVs with means $\mathbb{E}[X] = \mathbb{E}[Y] = 0$, second moments $\mathbb{E}[X^2] = \mathbb{E}[Y^2] = 1$, and $\mathbb{E}[XY] = 0.5$. We want to estimate Y using a linear function of X given by

$$\widehat{Y} = uX + v$$

where u and v are constants to be designed. Compute the values of u and v which would make the mean squared error given by

$$g(u, v) = \mathbb{E}[(\widehat{Y} - Y)^2].$$

as small as possible, i.e., we want to minimize g(u, v) with respect to variables u and v. Useful algebraic identity: $(a + b + c)^2 = a^2 + b^2 + c^2 + 2ab + 2bc + 2ca$.

Solution:
$$u = 0.5, v = 0$$

$$g(u, v) = \mathbb{E}[(uX + v - Y)^2]$$

= $u^2 \mathbb{E}[X^2] + v^2 + \mathbb{E}[Y^2] + 2uv\mathbb{E}[X] - 2v\mathbb{E}[Y] - 2u\mathbb{E}[XY]$
= $u^2 + v^2 + 1 - u$
= $(u - 0.5)^2 + v^2 + 0.75$

Thus, g(u, v) is minimized if u = 0.5 and v = 0 and the minimum mean squared error is 0.75.